# Peer Effects in Young Adults' Marital Decisions ${ }^{1}$ 

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#### Abstract

This paper studies peer group effects on marital decisions. We use data from Waves I and III of the National Longitudinal Study of Adolescent Health (Add Health). This database contains detailed information on adolescents' high school friends as well as their marital behavior later in life. We construct a balanced panel for the years 1995-2002 using the calendar of all past and current relationships of the respondents, which allows us to recover the marital status of each individual and of her friends at any given year. Hence, we are able to analyze how the marital transitions of individuals depend on the marital status of their friends. We use panel data, instrumental variables techniques, and exploit the timing of friendship formation to identify the effect of peers on marital decisions. Our results after controlling for various observable characteristics of individuals and their friends show that peer effects in marital decisions are significant. Robustness checks using former and placebo friends support our results, and indicate that actual peers do matter.


JEL classifications: A14, J12, J16, Z13
Keywords: peer effects, friends, marriage, cohabitation, marital decisions

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## 1 Introduction

Friends are an important part of individuals' life and constitute, together with the family, the social circle in which individuals develop. Friends often spend time together participating in the same activities (sports, school, etc), discussing about different topics and exchanging ideas. It is not then unrealistic to think that friends might affect each other's behavior through their opinions or through imitation. In fact, there is a large literature on peer group effects showing that friends actually affect, among others, the individual performance at school (Hoxby, 2000; Sacerdote, 2001; Calvó-Armengol, Patacchini and Zenou, 2009; Boucher, Bramoullé, Djebbari, and Fortin, 2010), obesity (Cohen-Cole and Fletcher, 2008), smoking habits (Gaviria and Rafael, 2001; Powell, Tauras and Ross, 2005; Fletcher, 2010: Card and Giuliano, 2011), and/or alcohol consumption (Clark and Lohéac, 2007; Fletscher, 2011), fertility (Kuziemko, 2006; Ciliberto, Miller, Nielsen, and Simonsen, 2010; Hensvik and Nillson, 2010), productivity (Falk and Ichino, 2006), the probability of finding a job (Topa, 2001; Calvó-Armengol and Jackson, 2004; Cappellari and Tatsiramos, 2010, Cingano and Rosolia, 2012), and the probability of engaging in criminal activities (Glaeser, Sacerdote and Scheinkman, 1996; Patacchini and Zenou, 2011). ${ }^{3}$

But what about marital decisions? Getting married or cohabiting is a decision that many young couples face. In order to make a decision, a potential couple might discuss with their parents, sibling or friends. If most friends of an individual are married, she may also want to get married in order, for example, to avoid being stigmatized by her friends. Likewise, if many of her friends are cohabiting she may also decide to do so. What the couple considers as a norm depends not only on the society as a whole (e.g. religion or tradition) but also on one's circle of close friends.

In this paper, we investigate whether the marital decisions of one's friends have any effect on one's own marital decisions. We use direct information on individuals' high school friends from the National Longitudinal Study of Adolescent Health (Add Health). We construct a balanced panel using the calendar of all past and current relationships of the respondents, which allows us to recover the marital status and other characteristics of each individual and of her friends at any given year. We

[^1]motivate our empirical exercise with a model of conformism and our results show that conformism might be the key mechanism behind the observed peer group effects. We find that when the percentage of married peers increases by $10 \%$ the probability of an individual getting married increases by 3.1 percentage points. The effect is statistically significant for females but not for males.

The biggest obstacle in identifying peer effects in marital decisions of the individuals has been data availability. In order to investigate the extent of peer group effects, Billari et al. (2007) use simulated data to show that social influence is the key driving force of the process of first marriage. Drewianka (1999 and 2003) uses data from PUMS and shows that a 10 percentage point increase in the fraction of persons aged 16-44 in a geographical area who are single leads to a decrease in individual's propensity to marry of an order of 1.5-2.0 percentage points. Moreover, he finds evidence that social effects operate through markets (search externality) and not directly through stigma or role modelling. Not only initiation but also termination of marriage might be influenced by peers. McDermott, Folwer, and Christakis (2009) show that divorce can spread between friends, siblings and coworkers.

Let us first discuss briefly the rapid changes in marital behavior that took place during the last decades. As Figure 1 shows, the marriage rate in the US has fallen drastically. Similar changes have been observed also in other developed countries.

Figure 1. Marriage Rate in the U.S.


Source: National Center for Family \& Mariage Research

Many studies have tried to identify the factors behind the drop in the marriage rate (declining gender gap, Becker 1981; contraceptive pill, Goldin \& Katz, 2002; household production technology, Greenwood \& Guner, 2009 to name a few). These forces are
likely to be amplified if there exist peer group effects that create a social multiplier. As a result, the effect of family-friendly policies, tax reforms, divorce laws, etc. will be augmented. Part of this drop is also due to the increase in the median age at first marriage (Figure 2). Individuals nowadays get married at an older age than what they used to do in the past. Hence, also the timing of marriage may be contagious in the sense that individuals decide to get married after observing that one of their peers got married.

Figure 2. Estimated Median Age at First Marriage in the U.S.


Source: U.S. Census Bureau

Figure 3. Cohabitation Rate in the U.S.


Source: U.S. Census Bureau

Cohabitation is a more recent phenomenon that is becoming more and more popular especially among young couples (Bumpass \& Lu, 2000). This upward trend (Figure 3), that is also present in western European countries, has been attributed to economic factors like the gender wage gap and the household production technology (Adamopoulou,
2010), female labor force participation and tertiary education (Kalmjin, 2007) or tax reforms (Leturcq, 2009). In addition to these factors, there might be an imitation effect (peer effect) at work that self-enforced the increase of cohabitation. Our aim is to identify peer effects in the decision to get married or cohabit as well as in the timing of these actions. In the next section we use a model of conformism in order to study a possible way friends can influence an individual's marital decisions. The model motivates the empirical analysis that follows.

## 2 Model

The model is based on Patacchini and Zenou (2011), who study juvenile delinquency using a model of conformism. The key element of the model is the notion of conformism, i.e., quoting the authors description, 'the idea that the easiest and hence best life is attained by doing one's very best to blend in with one's surroundings and to do nothing eccentric or out of the ordinary in any way'. Conformism might also be important for young adults when they decide whether to get married or cohabit. We first define the network structure of agents' friendships, and we then describe the preferences of the agents.

There is a finite number of agents $N=\{1, \ldots, n\}$. Let $g$ denote a particular network. We use the n-square adjacency matrix $G$ of a network $g$ to keep track of the direct connections in this network (see Jackson, 2008). Two agents $i$ and $j$ are directly connected in the network $g$ if and only if $g_{i j}=1$. We set $g_{i i}=0$, i.e. the agent cannot be a friend of herself. The set of direct connections of agent $i$ is $N_{i}(g)=\left\{j \neq i \mid g_{i j}=1\right\}$, which is of size $g_{i}=\sum_{j=1}^{n} g_{i j}$. In general $N_{i}(g) \neq N_{j}(g)$, unless the network is complete and everybody is a friend of everybody.

Each agent decides whether to stay single and just date with a partner, cohabit or get married. We assume, therefore, that there are many different degrees of formality that the relationship can take ranging from very informal (dating) to very formal (getting married). We denote the formality of the relationship by $f_{i}$. We then define the average formality of the relationships of $i$ 's friends as $\overline{f_{i}}(g)=\frac{1}{g_{i}} \sum_{j=1}^{n} g_{i j} f_{j}$

Each agent selects a degree of formality $f_{i} \geq 0$ for her relationship and receives a
payoff $u\left(f_{i}, \overline{f_{i}}\right)$ given by the utility function

$$
u_{i}\left(f_{i}, \overline{f_{i}}\right)=a+b_{i} f_{i}-\pi f_{i} \phi-c f_{i}^{2}-d\left(f_{i}-\overline{f_{i}}\right)^{2}
$$

with $a, c, d>0$, and $b_{i}>0, \forall i$.
There is a benefit from formalizing the relationship, which is given by the term $a+b_{i} f_{i}$. The agents are ex ante heterogeneous with respect to $b_{i}$. The parameter $b_{i}$ is assumed to be deterministic and observable by all agents in the network and it represents observable characteristics of individual $i$ (e.g., gender, race, age, education, religion etc.) and to the observable average characteristics of individual $i$ 's friends (contextual effects). More specifically,

$$
b_{i}(x)=\sum_{m=1}^{M} \beta_{m} x_{i}^{m}+\frac{1}{g_{i}} \sum_{m=1}^{M} \sum_{j=1}^{n} \theta_{m} g_{i j} x_{j}^{m},
$$

where $x_{i}^{m}$ are observable characteristics of individual $i$, the term $\frac{1}{g_{i}} \sum_{m=1}^{M} \sum_{j=1}^{n} g_{i j} x_{j}^{m}$ captures the contextual effects, and $\beta_{m}, \theta_{m}$ are parameters.

There is also a cost of formalizing the relationship, which is given by the term $-\pi f_{i} \phi-c f_{i}^{2}$. The parameter $\pi$ is the probability that the relationship ends and the parameter $\phi$ is the cost of ending the relationship. The cost of ending a relationship increases as the formality increases, i.e. it is more costly to separate if one is married than if one is cohabiting. Likewise, it is more costly to separate if one is cohabiting with a partner than if one is just dating this partner. The term $-c f_{i}^{2}$ is needed so as the cost function to be convex. Transiting from cohabitation to marriage is a more complicated procedure than transiting from dating to cohabitation.

The last term in the utility function, $-d\left(f_{i}-\bar{f}_{i}\right)^{2}$, reflects the influence of friends' behavior on own action. Each agent tries to minimize the distance between herself and her group of friends. The agent loses utility from failing to conform to others. Parameter $d$ represents the taste for conformity.

In this framework there exists a unique Nash equilibrium (Patacchini and Zenou, 2011) where each individual chooses the optimal formality of relationship $f_{i}^{*}$

$$
f_{i}^{*}=\frac{d}{c+d} \overline{f_{i}}+\frac{b_{i}}{2(c+d)}-\frac{\pi \phi}{2(c+d)},
$$

which is increasing in $\overline{f_{i}}$. In other words, the more formal the relationships of one's friends are, the more the individual will formalize her own relationship.

In the next section we test this result empirically and we try to figure out whether the percentage of individuals' married and cohabiting friends has any effect on individuals' decisions to enter into marriage or cohabitation. We also discuss other possible mechanisms that may drive the peer effect in marital decisions and provide evidence that support conformism as the main mechanism.

## 3 Data

Add Health is a longitudinal study of a nationally representative sample of adolescents in grades 7-12 in the United States during the 1994-95 school year. ${ }^{4}$ In Wave I the study started with an in-school questionnaire that was administered to more than 90,000 students from 80 high schools and 52 middle schools. A subsample of them (around 20,000 ) were also asked to complete in-home interviews and were followed in subsequent waves (II, III, and IV). The last wave was conducted in 2008, when the sample was aged 24-32. Adolescents had to answer questions about their family background, school performance, area of residence, tobacco and alcohol consumption, criminal activities as well as about sexual behavior (contraception, pregnancy, HIV and STD). In Wave I adolescents' mothers were also interviewed, and as a result, we can obtain information on their characteristics as well. However, mothers were not interviewed in the subsequent waves so it is not possible to update this information.

Wave III in-home interviews took place in years 2001 and 2002 and were completed by around 15,000 respondents aged 18 -28. In Wave III the respondents had to list all their current and previous sexual relationships ( $82 \%$ non missing responses) providing detailed information on the starting and ending date, whether they cohabited and how long, when they got married etc. Using this information we create a balanced panel for the years 1995-2002. For example, if a respondent listed a relationship with a partner for the years 2000-2002 with whom she started cohabiting in 2001 and she got married

[^2]in 2002, we will consider her single for the year 2000, cohabiting in 2001, and married in 2002. If the respondent had more than one relationship in a given year we keep the one with the longest overall duration. The procedure is similar to the one in Xie et al. (2003), and Raley et al. (2007) that analyze the determinants of marital transitions.

In Wave I, data collectors assigned an identification number to each student and provided a list of all students to the respondents in order to identify their friends. Respondents were allowed to list up to five male friends and up to five female friends. We treat two students as friends if at least one of the two has identified the other as his/her friend. As long as their nominated friends were also interviewed (i.e. they were part of the random subsample who completed the in-home survey), one can construct for each respondent a set of friends with detailed Add Health information. Given that the data represent a subsample of students within schools, not all nominated friends are interviewed and as a result, the measures of friends' characteristics will be imperfect. However, since the sampling scheme was random within grades, and most friends were in the same grade, the measures should be on average correct.

In Wave III, when the respondents were between ages 18 and 28, those who were in grades 7 and 8 in Wave I (two youngest grades), were interviewed about their current and former friends. In particular, data collectors presented each respondent with a list of 10 names and asked if any of them is currently or used to be their friend. For former friends, they also asked to state when the friendship ended and why. This list was created, based on an algorithm, using information from club membership and other school activities. As a result, it is possible for some respondents to identify current as well as former friends from this list. Furthermore, for former friends it is possible to know the year that the friendship has ended. By matching the identification numbers of friends to respondents' identification numbers we obtain information on the characteristics of nominated friends. In this way we know at any given year the marital status of the respondent and the marital status of his/her friends. Our final sample consists of 2,644 respondents with non missing relationship history that have at least one friend with non missing relationship history as well. The descriptive statistics of the individuals in our final sample are similar to the ones of all the individuals interviewed in wave III, ensuring that the final sample is still representative (see Table A1).

## 4 Empirical Strategy

Individual behavior may move conjointly with average peer group behavior for three different reasons. i) Endogenous effects; the behavior of the individual is causally influenced by the behavior of the group. This is the peer effect that we are trying to estimate. ii) Contextual effects; the behavior of the individual is influenced by the characteristics of the group. For example an individual might decide to get married because her friends are very religious independently from whether the friends are married or not. iii) Correlated effects; the individual and the group behave in the same way due to similar environments that are unobserved or due to endogenous friendship formation/sorting. This arises either from the fact that both the individual and her friends are subject to common unobserved shocks or because the individual selects friends who are similar to her.

Manski (1993) shows that identifying the endogenous and the contextual effects separately in a reduced form linear model is not possible. This is called the reflection problem and it is due to the fact that group behavior is by definition the aggregation of individual behavior. Solutions that have been proposed in order to solve the reflection problem consist of using instrumental variables techniques, or using panel data (see Bramoullé, Djebbari, and Fortin, 2009; Boucher et al., 2010). Instruments are used in order to generate variation in peer behavior that is independent from individual behavior. Examples of identification strategies with instrumental variables include Ciliberto et al. (2010) that use the fertility of the siblings of one's colleagues as an instrument for the fertility of one's colleagues, and Fletscher (2011) that uses the alcohol consumption of the parents of one's classmates as an instrument for the alcohol consumption of one's classmates. The basic idea is that siblings or parents of peers affect the behavior of the peers but have no independent effect on the respondent's behavior. De Giorgi, Pellizzari, and Redaelli (2010), and Pattachini and Zenou (2011) exploit the information about the whole network of friendships and instrument the behavior of the respondent's friends with the characteristics of friends of friends who are not directly linked with the respondent. With panel data one can focus on changes in the behavior over time in order to deal with the reflection problem. In this way, the reflection problem will disappear since the characteristics of the peers are already determined at the time that the change in individual behavior (transition
into employment, having a child, getting married etc.) takes place. Clark and Lohéac (2007) use panel data from Waves I and II of AddHealth to examine risky behavior (the consumption of tobacco, alcohol and marijuana). Kuziemko (2006) uses panel data in order to show that fertility is contagious among siblings. Cappellari and Tatsiramos (2011) use panel data to show that employed friends increase the probability of transition into employment. In an alternative identification strategy they consider the effect of the respondent on friends' transitions and instrument the respondent's employment status with the health status.

We instrument the percentage of married and cohabiting peers using the contextual variables. We thus assume that there is no direct effect of friends' characteristics on respondents' decisions and use friends' characteristics as instruments for their marital behavior. This procedure is common in the literature (e.g. Gaviria and Raphael, 2001; Powell et al., 2005).

What about correlated effects? One might worry that people make new friends as they get married, often through their spouse. Hence, it is normal for married people to make new friends who are also married. In this case endogeneity would be a serious problem in identifying the peer effects. In the current analysis we consider friends since high school and we have information about friendship dynamics. This solves part of the endogenous friendship formation in later years. Moreover it is not very likely that adolescents selected friends in high school according to characteristics that determined their marital behavior afterwards.

We use a panel data fixed effects estimator in order to deal with the correlated effect. Assuming that any correlation between the behavior of the peers and individual unobserved traits is due to traits that do not vary over time a panel data fixed effect estimator can deal with the correlated effect. Other studies that use a fixed effect estimator are Kuziemko (2006) and Cappellari and Tatsiramos (2011). Further robustness/falsification tests using placebo peer groups in the spirit of Fletscher (2011), and Hensvik and Nillson (2010) show that the peer effect is not due to selection.

## 5 Regression analysis

The benchmark regression is

$$
f_{i t}=\overbrace{\xi^{M} \overline{f_{i t}^{M}}+\xi^{C} \overline{f_{i t}^{C}}}^{\begin{array}{c}
\text { endogenous } \\
\text { effects }
\end{array}}+\underbrace{\sum_{m=1}^{M} \beta_{m} x_{i t}^{m}}_{\begin{array}{c}
\text { individual characteristics } \\
\text { (gender, age, race, etc) }
\end{array}}+\underbrace{\frac{1}{g_{i}} \sum_{m=1}^{M} \sum_{j=1}^{n} \theta_{m} g_{i j} x_{j t}^{m}}_{\begin{array}{c}
\text { average peer characteristics } \\
\text { (contextual effects) }
\end{array}}+y_{t}+\varepsilon_{i t}
$$

where $f_{i t}$ is a binary variable that takes the value 1 if an individual gets married (i.e. the individual was not married at $t-1$ and gets married at $t$ ), and 0 otherwise, $\overline{f_{i t}^{M}}$, $\overline{f_{i t}^{C}}$ are the percentages of married and cohabiting peers, $\xi^{M}$ and $\xi^{C}$ are the coefficients of interest, i.e. the peer effect that we are trying to estimate, $x_{i t}^{m}$ are the individual characteristics of the respondents ( $m$ variables that include gender, age, education, race, religiosity, beauty, relationship duration, out of wedlock births, mother's education, mother's age at first marriage, whether the mother was married in Wave I, and whether the mother has ever cohabited), $\frac{1}{g_{i}} \sum_{j=1}^{n} g_{i j} x_{j t}^{m}$ are the average individual and maternal characteristics of $i^{\prime} s n$ peers (contextual variables), i.e. the percentage of female peers, average education, percentage of African American peers, average religiosity, average beauty, average relationship duration, percentage of peers with out of wedlock births, average maternal education, average mother's age at first marriage, percentage of individuals whose mother was married in Wave I, and percentage of individuals whose mother has ever cohabited. ${ }^{5} y_{t}$ are year dummies.

### 5.1 Wave I nominations

We first examine the determinants of the transition into first marriage using the friends nominations from Wave I. Here, we assume that friendships have lasted after high school up to Wave III (i.e. for 7 years). This assumption will be relaxed afterwards using the updated information from Wave III (only for the subsample that this information is available). Table 1 shows the descriptive statistics for a total of 2,644 respondents with non missing own and peer relationship information. Around $67 \%$ of the respondents have one friend, $14 \%$ have 2 friends, $6 \%$ have 3 friends, $5 \%$ have 4

[^3]friends, $3 \%$ have 5 friends and less than $3 \%$ have $6-8$ friends. ${ }^{6}$
We start our analysis with a linear probability model (Table 2, column 1). The dependent variable takes the value 1 if someone who was not married in the previous year gets married in the current year, and the value 0 otherwise. The variables of interest are the ratio of each individual's friends that are cohabiting and the ratio of friends that are married. We include as regressors the characteristics of the individuals, such as age, gender, race, education, religiosity, and a measure of beauty (the interviewer had to assess the physical attractiveness of the respondent). All variables are explained in the appendix. We also include whether a respondent had an out-of-wedlock birth in the past as this might affect the probability of getting married. We account for maternal characteristics, such as mother's marital status at Wave I, mother's education, mother's age at first marriage and whether the mother has ever cohabited. We include the duration of the relationship which also acts as a control for being in a relationship (when someone is not in a relationship, relationship duration will be zero). All these are variables commonly used in the literature when studying the determinants of marital behavior (see Raley et al., 2007). Finally, we include year dummies in all specifications. We use the appropriate weights and robust standard errors clustered at the school level. In this specification we also include contextual variables, i.e. the average individual characteristics of the peers. The percentage of married peers has a statistically significant effect on the transition into marriage. If the percentage of married peers increases by $10 \%$ the probability of an individual getting married increases by 2.3 percentage points. The percentage of cohabiting peers does not seem to matter.

We then perform 2SLS in order to improve the identification (Table 2, column 2, see Tables A3 and A4 for the 1st stage regressions). Following the literature (Gaviria and Raphael, 2001; Powell et al., 2005) we assume that the contextual variables do not have any effect on individual behavior, i.e. $\theta_{m}=0$ (indeed their effect was statistically insignificant in the OLS) and we exclude them from the regression. Instead, we use these contextual variables as an instrument for the percentage of married and

[^4]cohabiting peers. ${ }^{7}$ The F statistic of the excluded instruments in the 1st stage is larger than 10 (18.43 for the percentage of married peers and 23.25 for the percentage of cohabiting peers) indicating that the instruments are not weak. The Hansen J statistic does not reject the hypothesis of the validity of the instruments. The effect of married peers remains statistically significant and its magnitude increases. In particular, if the percentage of married peers increases by $10 \%$ the probability of an individual getting married increases by 6.2 percentage points.

Alternatively, we perform a panel data fixed effect estimation that also can deal with the identification issues (Table 3, column 1). In this specification we include only time varying variables (age, education, out of wedlock births and the duration of the relationship). We also control for the average education and relationship duration of the peers. We cannot include parental characteristics because we have information only for Wave I (and hence no time variation). The peer effect remains significant but decreases in magnitude.

Lastly, we extend our network to the friends of friends who are not directly connected with the respondents (Table 3, column 2). The percentage of married friends of friends has a similar effect although not statistically significant. Hence, there is no clear evidence of spill-over effects between individuals that are only indirectly connected with each other.

We also perform the analysis for girls and boys separately to see whether there are any gender differences with respect to the magnitude, significance or the direction of the effect. The peer effect on girls is positive, and statistically significant (Table 4, columns 1 and 2), while the peer effect on boys (Table 4, columns 3 and 4) is not statistically significant. This result might reflect a stigma towards unmarried females that is stronger than towards males. On the other hand, this gender difference might just reflect the fact that girls have more female friends than boys, and females get married at an earlier age than males. More specifically in our sample almost $21 \%$ of girls got married by 2002 while this percentage falls to $12 \%$ for boys. Around $56 \%$ of girls have only female friends and $45 \%$ of boys have only male friends.

[^5]
### 5.1.1 Same gender friends

One might worry that respondents nominate as friends individuals of different gender with whom they have a sexual relationship. ${ }^{8}$ In this case, the peer effect would be spurious. Suppose that a male respondent nominates as a friend a girl with whom he has a relationship and eventually he gets married. Hence, the percentage of married peers increases and at the same time he transits into marriage. A situation like this would be mistakenly considered as peer effect although in reality the respondent and his peer had married each other. For this reason we conduct the same analysis using same-gender friends only. The results remain almost unchanged (Table 5) indicating that there should be no concern about marriages among peers.

### 5.1.2 Differential peer effect

But who are the ones who are influenced by their peers? Are they all the individuals or only some particular groups? In order to answer this question we analyze separately different groups of individuals with respect to religiosity and race. Marital behavior differs substantially between religious and non religious individuals, white and African Americans, hence the peer effect might also differ. Indeed, when we repeat the analysis for different groups we find that the peer effect vanishes for the non-religious ones (defined as those who have never attended religious services in the past 12 months), and it becomes stronger for the religious ones (Table 6, columns 1 and 2). Moreover, the peer effect is present only for less educated individuals and not for those that they continue with their studies after highschool (Table 6, columns 3 and 4). More religious and less educated individuals are the ones that get married early, and therefore the effect comes from these particular groups of people.

### 5.1.3 Cohabitation

Next, we conduct the same analysis for the transition into cohabitation (Table 7). In this case the dependent variable takes the value 1 if an individual was not cohabiting at $t-1$ and starts cohabiting at $t$, and is zero otherwise. Our variables of interest are again the percentages of married and cohabiting peers. On the one hand, we find no

[^6]statistically significant effect of the percentage of cohabiting peers. This means that if an individual has many cohabiting peers this will not increase her probability of cohabiting. On the other hand, there is a negative effect of the percentage of married peers on the decision to cohabit. We interpret this as evidence in favor of conformism with respect to marriage. If the percentage of married peers increases by $10 \%$ the probability of entering cohabitation decreases by 5.5 percentage points. Hence, having married peers acts as a deterrent to cohabitation.

### 5.1.4 Timing of the transition

As we discussed in the introduction, the age of first marriage has increased during the last decades. Hence, it might be the case that the timing of marriages is contagious. To study this we check whether the transition into marriage in a given year is affected from any peer who got married in the previous year or from any peer who started cohabiting in the previous year. The results in Table 8 suggest that the peer effect in the timing of marriage is significant and similar in magnitude as the peer effect in the decision to get married. Moreover, also in the case that a peer started cohabiting in the previous year the respondent's probability of getting married this year goes up (but less than if the peer got married). There is no peer effect in the timing of cohabitation (Table 9). There is evidence of a negative effect on the timing of cohabitation if some peer got married in the last year, supporting the conformistic behavior that is entailed with marriage.

### 5.2 Mechanism

From the analysis so far, there is evidence of a conformistic behavior with respect to marriage. However, there might be alternative underlying mechanisms, such as leisure complementarities or search externalities. A mechanism of leisure complementarities will drive individuals into marriage in order to share common interests with their married peers. The lifestyle of married people is different from the one of singles. As a result, a single individual with many married peers might decide to get married in order to be able to spend time with them doing similar activities. We use the geographical proximity of friends in order to see whether this mechanism is at work. Small geographical distance between friends facilitates communication and encourages
them to enjoy leisure together. We use information on the county of residence of the peers and the respondent in Wave III (Table 10). We focus on respondents that have both a peer that resides in the same county as them and a peer that resides in different county. This reduces the sample size. The coefficient of the percentage of same-county married friends is not statistically significant. Hence, a mechanism of leisure complementarities does not seem to drive the peer effect on marriage.

The alternative mechanism of search externalities would induce people into marriage through competition on available partners. More specifically, if most of the friends of an individual are married or cohabit, the set of available partners shrinks and this might make marriage more urgent. However, we have found that the percentage of cohabiting peers does not have any significant effect. We would expect this not to be the case if search externalities were at work. ${ }^{9}$

### 5.3 Wave III nominations: current friends

Up to now we were using friendship nominations from Wave I in order to define the peer group of the respondents. In other words, we were assuming that friendships have lasted throughout the years until Wave III. In this section we are going to relax this assumption by updating the friendship information. As we already mentioned, for the respondents of Wave III who were in the 7th or 8th grade at Wave I, an algorithm, based on clubs and activities from previous waves, was used to select 10 names of students who also attended the same school. These respondents were then asked to identify whether or not they were currently or had been previously friends with each of the 10 listed names. We then perform the analysis by using current friends as the peer group of reference. Our sample consists of 1,065 respondents who identified at least one current friend and have non missing own and peer relationship history. Table 11 shows the descriptive statistics for these respondents. ${ }^{10}$ To our knowledge this is the first study that uses the friendship information from Wave III of the Add Health survey. One reason for this might be the small sample size. This is not an issue in our case thanks to the retrospective panel of relationships that we have constructed.

[^7]We get larger estimates when we repeat the analysis using only the friends that the respondents have identified as current ones (Table 12, columns 1 and 2 ). The effect is almost double in magnitude ( 0.057 compared to 0.031 in the fixed effects estimation and 0.133 compared to 0.062 in the 2SLS). This is not surprising given that current friends are expected to exert bigger influence than the whole set of high school friends that contains both current and former friends. There is no clear evidence of spillover effects from current friends of current friends either (Table 12, column 3). The results in Table 13 regarding the transition into cohabitation provide us with a further confirmation of the conformistic mechanism based on the big negative effect of current married peers.

## 6 Robustness

At this point one may think that it is natural to find a positive correlation among individuals that went to the same school and share many common characteristics and thus may doubt about the causality of the peer effect. In order to convince the reader, we perform robustness checks using different groups of peers, namely ghost and placebo friends. The idea behind it is that if the peer effect is spurious it must show up also when considering as peer group of reference individuals with similar characteristics as the real friends who nevertheless are not connected to the respondents.

### 6.1 Ghost friends

The first robustness check uses "ghost" friends as the peer group of reference. We define ghost friends as follows. There are respondents who indicated that they had been previously friends (but not anymore) with some of the 10 names that they were provided with in Wave III. Moreover, we have information on the exact year that the respondent last saw the former friend in person, talked with her on the telephone, or exchanged email. We can thus consider the effect of ghost friends, i.e. the effect of former friends in the years after the friendship has ended. We expect that ghost friends should not have any effect on the decisions of the individuals. However, there might be concerns regarding the reasons that the friendship has ended. ${ }^{11}$ If the friendship

[^8]has ended due to the fact that former friends got married, ghost friends would not be adequate for our robustness check. In Table 13 we display the descriptive statistics of ghost friends in comparison with the ones of current friends. We do not observe a bigger tendency towards marriage for ghost friends compared to current friends.

We perform the same analysis using ghost friends instead of current friends. The percentage of married ghost friends does not have a significant effect on the transition of individuals into marriage although the magnitude of the effect is similar to the one of real friends (Table 14). ${ }^{12}$ This result also alleviates the concern that individuals may chose with whom to remain friends in an endogenous way.

### 6.2 Placebo friends

A further robustness check consists of using placebo friends as the peer group of reference. Remember that in Wave III an algorithm, based on clubs and activities from previous waves, was used to select 10 names of students who also attended the same school. In certain cases the respondents indicated that they did not know some of the 10 names. The unidentified names correspond to individuals that could have been potentially friends with the respondent given that the 10 names were not random, but the algorithm selected them among students of the same school who were doing similar activities with the respondent. Thus, we can exploit this feature of the algorithm and define these unidentified individuals as placebo friends. Table 15 demonstrates that the characteristics of placebo and real friends are similar. Not surprisingly, when we perform the robustness check placebo friends do not have any significant effect either (Table 16). It is thus actual peers, and not just students from the same school that do matter for the decisions of the respondents. This robustness check is supportive of a causal interpretation of the effect of real friends.

### 6.3 Friends that enter into marriage/cohabitation the year after

The last robustness check concerns the timing of the transition into marriage. As we saw in Table 8, if any peer got married in the previous year, this would affect the away, the friend changed, you changed, the friend died.
${ }^{12}$ We may lose statistical significance also due to the small sample size.
transition of the respondent into marriage in the current year. However, we expect that if any peer gets married next year, this will not have any effect on the transition of the respondent into marriage in the current year. Indeed, this is the case as Table 17 shows. Hence, the timing of marriage is indeed contagious.

## 7 Conclusions

The analysis shows a positive and significant peer effect on the transition of singles into marriage. Increasing the proportion of married peers by $10 \%$ leads to an increase in individual's propensity to get married on the order of $0.3-1.3$ percentage points. The effect is present for girls, religious and white people. The fact that there is no significant effect of ghost and placebo friends indicates that real peers do matter. There does not seem to exist a peer effect on the transition into cohabitation. Instead, there is a negative effect of the percentage of married peers on the respondents' transition into cohabitation. This is an indication of a conformistic behavior with respect to marriage. There is no evidence of leisure complementarities or search externalities. The social multiplier has to be taken into account when analyzing the effect of family-friendly policies, tax reforms, divorce laws or other policies that may affect the incentives to get married.

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Table 1. Individual characteristics in $2002^{1,2}$ (Wave I nominations) Characteristic

| \% females | $55.27(0.013)$ |
| :--- | :---: |
| Mean Age | $22.42(0.186)$ |
| \% cohabiting | $19.24(0.014)$ |
| \% married | $16.94(0.017)$ |
| \% African American | $10.57(0.024)$ |
| \% with >high school education | $61.09(0.031)$ |
| Mean Religiosity (7-category scale) | $1.95(0.089)$ |
| Mean Beauty (5-category scale) | $3.57(0.032)$ |

${ }^{1}$ Individuals with non missing own and peers' relationship history
${ }^{2}$ Corrected for survey design

Table 2. Determinants of transition into marriage (Pooled OLS and 2SLS)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | Pooled OLS | 2 SLS |
| Definition of Peers | Nominated friends | Nominated friends |
| $\%$ married peers | $0.023^{* *}$ | $0.062^{* *}$ |
|  | $(0.011)$ | $(0.030)$ |
| $\%$ cohabiting peers | 0.004 | -0.051 |
|  | $(0.065)$ | $(0.045)$ |
| Individual characteristics | Yes | Yes |
| Parental characteristics | Yes | Yes |
| Contextual characteristics | Yes | Used as instruments |
| No of person-years | 15,709 | 14,662 |
| R $^{2}$ | 0.069 | 0.054 |
| F-statistic 1st stage | - | $18.43 ; 23.25$ |
| J statistic p value | - | 0.6214 |

${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level), cross sectional weights used Individual characteristics: gender, age, education, race, religiosity, beauty, relationship duration, out of wedlock births, Parental characteristics: marital status at Wave I, mother's education, age at first marriage, whether ever cohabited, Excluded instruments: peers' education, religiosity, beauty relationship dur, out of wedlock births, $\%$ females, $\%$ African Americans, $\%$ with married mothers Year dummies included in all specifications

Table 3. Determinants of transition into marriage (Fixed effects)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | Linear Prob/FE | Linear Prob/FE |
| Definition of Peers | Nominated friends | Nominated friends of friends |
| $\%$ married peers | $0.031^{* *}$ | 0.029 |
|  | $(0.013)$ | $(0.024)$ |
| $\%$ cohabiting peers | 0.009 | 0.000 |
|  | $(0.007)$ | $(0.012)$ |
| Individual characteristics | Yes | Yes |
| Parental characteristics | No | No |
| Contextual characteristics | Yes | Yes |
| No of person-years | 19,629 | 10,364 |
| $\mathrm{R}^{2}$ | 0.053 | 0.052 |
| $* * *$ p $<0.01,{ }^{* *}$ p<0.05, * p<0.1 (robust s.e. clustered at school level), cross sectional weights used |  |  |
| Individual characteristics (time varying): age, education, out of wedlock births, relationship dur. |  |  |
| Contextual characteristics: average education, and average relationship duration |  |  |
| Year dummies included in all specifications |  |  |

Table 4. Girls' and boys' determinants of transition into marriage (2SLS and fixed effects)

|  | Girls |  | Boys |  |
| :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) |
| Specification | 2SLS | Linear Prob/FE | 2SLS | Linear Prob/FE |
| Definition of Peers | Nomin. friends | Nomin. friends | Nomin. friends | Nomin. friends |
| \% married peers | 0.099** | 0.032* | 0.039 | 0.029 |
|  | (0.047) | (0.017) | (0.038) | (0.022) |
| \% cohabiting peers | -0.097 | 0.005 | -0.028 | 0.015 |
|  | (0.066) | (0.011) | (0.064) | (0.010) |
| Individual char. | Yes | Yes | Yes | Yes |
| Parental char. | Yes | No | Yes | No |
| Contextual char. | Used as instrum. | Yes | Used as instrum. | Yes |
| No of person-years | 7,956 | 10,791 | 6,706 | 8,838 |
| $\mathrm{R}^{2}$ | 0.041 | 0.051 | 0.046 | 0.056 |
| F-statistic 1st stage | 12.30; 11.41 | - | 11.33; 6.12 | - |
| J statistic p value | 0.744 | - | 0.574 | - |

*** $\mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level), cross sectional weights used Control variables: see Table 2 for 2SLS specification, and Table 3 for FE specification

Table 5. Determinants of transition into marriage (same gender friends)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | 2 SLS | Linear Prob/FE |
| Definition of Peers | Same gender friends | Same gender friends |
| \% married peers | $0.060^{*}$ | $0.031^{* *}$ |
|  | $(0.033)$ | $(0.015)$ |
| \% cohabiting peers | -0.006 | 0.008 |
|  | $(0.042)$ | $(0.009)$ |
| Individual char. | Yes | Yes |
| Parental char. | Yes | No |
| Contextual char. | Used as instruments | 15,281 |
| No of person-years | 11,324 | 0.053 |
| $R^{2}$ | 0.055 | - |
| F-statistic 1st stage | $18.94 ; 16.75$ | - |
| J statistic | 0.666 | Yes |
| $* * *$ p<0.01, $* *$ p $<0.05, *$ p<0.1 (robust s.e. clustered at school level) |  |  |
| Control variables: see Table 2 for 2 SLS specification, and Table 3 for FE specification |  |  |

Table 6. Determinants of transition into marriage by characteristic

|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ |
| :--- | :---: | :---: | :---: | :---: |
| Specification | Linear Prob/FE | Linear Prob/FE | Linear Prob/FE | Linear Prob/FE |
| Definition of Peers | Religious | Non-religious | High school or less | More than high school |
| \% married peers | $0.037^{* * *}$ | 0.019 | $0.039^{*}$ | 0.017 |
|  | $(0.013)$ | $(0.023)$ | $(0.022)$ | $(0.021)$ |
| \% cohabiting peers | 0.010 | 0.010 | 0.007 | 0.010 |
|  | $(0.008)$ | $(0.013)$ | $(0.011)$ | $(0.013)$ |
| Individual char. | Yes | Yes | Yes | Yes |
| Parental char. | No | No | No | No |
| Contextual char. | Yes | Yes | Yes | Yes |
| No of person-years | 14,634 | 4,995 | 11,016 | 8,613 |
| $\mathrm{R}^{2}$ | 0.058 | 0.045 | 0.059 | 0.049 |
| $* * *$ p $<0.01$, ** p<0.05, ${ }^{*}$ p<0.1 (robust s.e. clustered at school level), cross sectional weights used |  |  |  |  |
| Individual characteristics (time varying): age, education, out of wedlock births, relationship dur. |  |  |  |  |
| Contextual characteristics: average education, and average relationship duration |  |  |  |  |
| Year dummies included in all specifications |  |  |  |  |

Table 7. Determinants of transition into cohabitation (2SLS and fixed effects)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | 2 SLS | Linear Prob/FE |
| Definition of Peers | Nominated friends | Nominated friends |
| $\%$ married peers | $-0.055^{*}$ | -0.003 |
|  | $(0.031)$ | $(0.018)$ |
| $\%$ cohabiting peers | 0.076 | 0.005 |
|  | $(0.049)$ | $(0.013)$ |
| Individual char. | Yes | Yes |
| Parental char. | Yes | No |
| Contextual char. | Used as instruments | Yes |
| No of person-years | 14,408 | 19,783 |
| $R^{2}$ | 0.029 | 0.020 |
| F-statistic 1st stage | $19.53 ; 18.86$ | - |
| J statistic | 0.717 | - |

${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level), cross sectional weights used Variables in the FE specification: age, education, rel. duration, peers' average education, rel. duration Variables in the 2SLS specification: age, race, gender, education, relationship dur, religiosity, beauty, all parental characteristics, Excluded instruments: peers' education, religiosity, beauty, rel. duration, out of wedlock births, \% females, \% African Americans, \% with ever cohabiting mother.

Year dummies in all specifications

Table 8. Determinants of the timing of the transition into marriage (Fixed effects)

|  | $(1)$ |
| :--- | :---: |
| Specification | Linear Prob/FE |
| Definition of Peers | Nominated friends |
| Any peer entered marriage in the previous year | $0.032^{*}$ |
|  | $(0.017)$ |
| Any peer entered cohabitation in the previous year | $0.013^{*}$ |
|  | $(0.007)$ |
| Individual characteristics | Yes |
| Parental characteristics | No |
| Contextual characteristics | Yes |
| No of person-years | 17,009 |
| $\mathrm{R}^{2}$ | 0.048 |
| $* * * \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level), cross sectional weights used |  |
| Control variables: see Table 3 |  |

Table 9. Determinants of the timing of the transition into cohabitation (Fixed effects)

|  | $(1)$ |
| :--- | :---: |
| Specification | Linear Prob/FE |
| Definition of Peers | Nominated friends |
| Any peer entered marriage in the previous year | $-0.017^{*}$ |
|  | $(0.010)$ |
| Any peer entered cohabitation in the previous year | 0.0003 |
| Individual characteristics | $(0.012)$ |
| Parental characteristics | Yes |
| Contextual characteristics | No |
| No of person-years | Yes |
| $\mathrm{R}^{2}$ | 17,170 |
| $* * *$ p<0.01, ${ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level), cross sectional weights used |  |
| Control variables: see Table 3 | 0.015 |

Table 10. Transition into marriage and geographical proximity

|  | Linear Prob/FE |
| :--- | :---: |
| Specification | Nominated friends |
| Definition of Peers | $0.122^{* *}$ |
| \% married peers | $(0.046)$ |
|  | -0.005 |
| \% same county married peers | $(0.034)$ |
| \% cohabiting peers | 0.043 |
|  | $(0.040)$ |
| \% same county cohabiting peers | -0.005 |
| Individual characteristics | $(0.018)$ |
| Parental characteristics | Yes |
| Contextual characteristics | No |
| No of person-years | Yes |
| $R^{2}$ | 3,352 |
| $* *$ p<0.01, ** p<0.05, * p<0.1 (robust s.e. clustered at school level), cross sectional weights used |  |
| Control variables: see Table 3 | 0.064 |

Table 11. Individual characteristics in $2002^{1,2}$ (Wave III nominations)

## Characteristic

| \% females | $49.37(0.018)$ |
| :--- | :---: |
| Mean Age | $20.61(0.036)$ |
| \% cohabiting | $19.20(0.015)$ |
| \% married | $7.27(0.012)$ |
| \% African American | $12.71(0.028)$ |
| \% with >high school education | $50.43(0.026)$ |
| Mean Religiosity (7-category scale) | $2.19(0.106)$ |
| Mean Beauty (5-category scale) | $3.56(0.047)$ |

${ }^{1}$ Individuals with non missing own and peers' relationship history
${ }^{2}$ Corrected for survey design

Table 12. Determinants of transition into marriage (current friends)

|  | $(1)$ | $(2)$ | $(3)$ |
| :--- | :---: | :---: | :---: |
| Specification | 2 SLS | Linear Prob/FE | Linear Prob/FE |
| Definition of Peers | Current friends | Current friends | Current friends of friends |
| \% married peers | $0.133^{* *}$ | $0.057^{*}$ | 0.058 |
| \% cohabiting peers | $(0.053)$ | $(0.029)$ | $(0.040)$ |
|  | -0.043 | 0.013 | 0.017 |
| Individual char. | $(0.034)$ | $(0.013)$ | $(0.014)$ |
| Parental char. | Yes | Yes | Yes |
| Contextual char. | Used as instruments | No | No |
| No of person-years | 6,637 | 8,252 | Yes |
| $R^{2}$ | 0.012 | 0.031 | 2,980 |
| F-statistic 1st stage | $4.95 ; 10.53$ | - | 0.021 |
| J statistic | 0.897 | - | - |

${ }^{* * *} \mathrm{p}<0.01,{ }^{* *} \mathrm{p}<0.05,^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level)
Control variables: see Table 2 for 2SLS specification, and Table 3 for FE specification

Table 13. Determinants of transition into cohabitation (current friends)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | 2 SLS | Linear Prob/FE |
| Definition of Peers | Current friends | Current friends |
| \% married peers | $-0.171^{*}$ | $-0.056^{* *}$ |
|  | $(0.099)$ | $(0.021)$ |
| \% cohabiting peers | 0.016 | 0.001 |
|  | $(0.063)$ | $(0.025)$ |
| Individual char. | Yes | Yes |
| Parental char. | Yes | Yes |
| Contextual char. | Used as instruments | 8,081 |
| No of person-years | 6,501 | 0.041 |
| $R^{2}$ | 0.058 | - |
| F-statistic 1st stage | $5.45 ; 8.16$ | - |
| J statistic | 0.555 |  |
| $* * *$ p $<0.01, * * \mathrm{p}<0.05, *$ | p $<0.1$ (robust s.e. clustered at school level) |  |
| Control variables: see Table 7 |  |  |

Table 14. Real vs ghost friends' characteristics in $2002^{1,2}$

| Characteristic | Real friends | Ghost friends |
| :--- | :--- | :--- |
| \% married | 7.27 | 8.47 |
| \% cohabiting | 19.20 | 20.66 |
| \% females | 49.37 | 60.22 |
| Mean Age | 20.61 | 20.63 |
| \% African American | 12.71 | 15.07 |
| \% with > high school education | 50.43 | 54.18 |
| Religiosity (5-category scale) | 2.19 | 2.06 |
| Beauty (5-category scale) | 3.56 | 3.61 |

${ }^{1}$ Individuals with non missing relationship history
${ }^{2}$ Corrected for survey design

Table 15. Determinants of transition into marriage (ghost friends)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | 2 SLS | Linear Prob/FE |
| Definition of Peers | Ghost friends | Ghost friends |
| \% married peers | -0.057 | 0.035 |
|  | $(0.102)$ | $(0.029)$ |
| \% cohabiting peers | 0.054 | -0.006 |
|  | $(0.071)$ | $(0.014)$ |
| Individual char. | Yes | Yes |
| Parental char. | Yes | Yes |
| Contextual char. | Used as instruments | 3,452 |
| No of person-years | 2,897 | 0.043 |
| $R^{2}$ | 0.067 | - |
| F-statistic 1st stage | $4.05 ; 11.07$ | - |
| J statistic | 0.914 |  |
| $* * *$ p<0.01, $* *$ p $<0.05, *$ | p $<0.1$ (robust s.e. clustered at school level) |  |
| Control variables: see Table 7 |  |  |

Table 16. Real vs placebo friends' characteristics in $2002^{1,2}$

| Characteristic | Real friends | Placebo friends |
| :--- | :--- | :--- |
| \% married | 7.27 | 9.71 |
| \% cohabiting | 19.20 | 21.16 |
| \% females | 49.37 | 60.53 |
| Mean Age | 20.61 | 20.72 |
| \% African American | 12.71 | 19.49 |
| \% with > high school education | 50.43 | 47.27 |
| Religiosity (5-category scale) | 2.19 | 2.09 |
| Beauty (5-category scale) | 3.56 | 3.55 |

[^9]Table 17. Determinants of transition into marriage (placebo friends)

|  | $(1)$ | $(2)$ |
| :--- | :---: | :---: |
| Specification | 2 SLS | Linear Prob/FE |
| Definition of Peers | Placebo friends | Placebo friends |
| \% married peers | 0.024 | 0.002 |
|  | $(0.066)$ | $(0.021)$ |
| \% cohabiting peers | 0.008 | 0.015 |
|  | $(0.507)$ | $(0.014)$ |
| Individual char. | Yes | Yes |
| Parental char. | Yes | Yes |
| Contextual char. | Used as instruments | 7,066 |
| No of person-years | 5,638 | 0.028 |
| $R^{2}$ | 0.044 | - |
| F-statistic 1st stage | $5.61 ; 16.23$ | - |
| J statistic | 0.652 |  |
| $* * *$ p<0.01, ** p<0.05, * p<0.1 (robust s.e. clustered at school level) |  |  |
| Control variables: see Table 7 |  |  |

Table 18. Determinants of the timing of the transition into marriage (Robustness)

|  | $(1)$ |
| :--- | :---: |
| Specification | Linear Prob/FE |
| Definition of Peers | Nominated friends |
| Any peer enters marriage in the year after | 0.012 |
|  | $(0.012)$ |
| Any peer enters cohabitation in the year after | -0.004 |
|  | $(0.006)$ |
| Individual characteristics | Yes |
| Parental characteristics | No |
| Contextual characteristics | Yes |
| No of person-years | 17,498 |
| $\mathrm{R}^{2}$ | 0.065 |
| $* * *$ p<0.01, $* * \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.1$ (robust s.e. clustered at school level), cross sectional weights used |  |
| Control variables: see Table 3 |  |

### 8.1 Appendix

Table A1. Sample's characteristics in $2002^{1}$

| Characteristic | All individuals ${ }^{2}$ | Sample $^{3}$ |
| :--- | :--- | :--- |
| \% married | 17.67 | 16.94 |
| \% cohabiting | 21.77 | 19.24 |
| \% females | 51.40 | 55.26 |
| Mean Age | 22.59 | 22.41 |
| \% African American | 15.04 | 10.56 |
| \% with > high school education | 53.05 | 61.09 |
| Religiosity (5-category scale) | 1.92 | 1.96 |
| Beauty (5-category scale) | 3.51 | 3.58 |
| N | 10,220 | 2,644 |
| ${ }^{1}$ Corrected for survey design |  |  |
| ${ }^{2}$ Individuals with non missing own relationship history |  |  |
| ${ }^{3}$ Individuals with non missing own and friends' relationship history |  |  |

Table A2. Definition of Variables


| Mother's education | binary | $\left\{\begin{array}{l}0 \text { if high school or less } \\ 1 \text { if more than high school }\end{array}\right.$ |
| :---: | :---: | :---: |
| Mother's age at first marriage | continuous | [13, 53] |
| Mother ever cohabited | binary | $\left\{\begin{array}{l}1 \text { if the mother has ever cohabited } \\ 0 \text { otherwise }\end{array}\right.$ |
| Out of wedlock births | binary | $\left\{\begin{array}{l} 1 \text { if birth before the } 9 \text { th month of marriage } \\ 0 \text { otherwise } \end{array}\right.$ |
| Relationship duration | continuous | in years ( $=0$ if not currently in a relationship) |
| Contextual |  | average of all characteristics |

Table A3. IV-Auxilliary Equation

| Instrumented: $\%$ married peers |  |  |
| :--- | :--- | ---: |
|  | Coefficient | Std. Error |
| Gender | -0.0041 | 0.0109 |
| Age | $0.0126^{* * *}$ | 0.0035 |
| Race | 0.0072 | 0.0438 |
| Education | -0.0013 | 0.0138 |
| Religiosity | 0.0026 | 0.0035 |
| Beauty | -0.0028 | 0.0045 |
| Mother married at wave I | $-0.0205^{*}$ | 0.0114 |
| Mother's age at first marriage | $-0.0035^{* * *}$ | 0.0010 |
| Mother's education | 0.0088 | 0.0076 |
| Mother ever cohabited | 0.0234 | 0.0156 |
| Out of wedlock birth | 0.0025 | 0.0218 |
| Relationship duration | -0.0016 | 0.0029 |
| y1995 | $0.0670^{* *}$ | 0.0269 |
| y1996 | $0.0466^{* *}$ | 0.0248 |
| y1997 | 0.0332 | 0.0234 |
| y1998 | 0.0236 | 0.0211 |
| y1999 | $0.0282^{*}$ | 0.0163 |
| y2000 | $0.0247^{* *}$ | 0.0104 |
| y2001 | $0.0283^{* * *}$ | 0.0060 |
| \% female peers | 0.0124 | 0.0125 |
| \% African American peers | $-0.0871^{* *}$ | 0.0428 |
| average rel. duration | $0.0465^{* * *}$ | 0.0060 |
| \% peers with out of wed birth | $0.2386^{* * *}$ | 0.0364 |
| \% peers with married mother | -0.0233 | 0.0175 |
| average education | $-0.0419^{*}$ | 0.0214 |
| average religiosity | $0.0151^{* * *}$ | 0.0041 |
| average beauty | $-0.0150^{*}$ | 0.0082 |
|  |  |  |

$\mathrm{R}^{2} \quad 0.278$
${ }^{* * *} \mathrm{p}<0.001,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.10$, rob s.e. clustered at school level
Cross sectional weights used
F test of excluded instruments: $\mathrm{F}(8,125)=18.43$, Prob $>\mathrm{F}=0.000$

Table A4. IV-Auxilliary Equation
Instrumented: \% cohabiting peers

|  | Coefficient | Std. Error |
| :--- | :--- | ---: |
| Gender | -0.0056 | 0.0101 |
| Age | $0.0114^{* * *}$ | 0.0036 |
| Race | $-0.0579^{* *}$ | 0.0259 |
| Education | $-0.0202^{*}$ | 0.0119 |
| Religiosity | -0.0020 | 0.0025 |
| Beauty | $0.0126^{* *}$ | 0.0052 |
| Mother married at wave I | $0.0328^{* *}$ | 0.0164 |
| Mother's age at first marriage | -0.0010 | 0.0012 |
| Mother's education | -0.0118 | 0.0079 |
| Mother ever cohabited | -0.0280 | 0.0188 |
| Out of wedlock birth | 0.0025 | 0.0230 |
| Relationship duration | 0.0038 | 0.0038 |
| y1995 | -0.0179 | 0.0257 |
| y1996 | -0.0092 | 0.0247 |
| y1997 | -0.0022 | 0.0238 |
| y1998 | -0.0009 | 0.0224 |
| y1999 | 0.0125 | 0.0208 |
| y2000 | 0.0054 | 0.0153 |
| y2001 | $0.0144^{* *}$ | 0.0068 |
| \% female peers | 0.0149 | 0.0121 |
| \% African American peers | 0.0354 | 0.0284 |
| average rel. duration | $0.0299^{* * *}$ | 0.0046 |
| \% peers with out of wed birth | $0.1150^{* * *}$ | 0.0382 |
| \% peers with married mother | $0.0263^{*}$ | 0.0151 |
| average education | $-0.0364^{* *}$ | 0.0141 |
| average religiosity | $-0.0142^{* * *}$ | 0.0023 |
| average beauty | -0.0127 | 0.0078 |
|  |  |  |

${ }^{* * *} \mathrm{p}<0.001,{ }^{* *} \mathrm{p}<0.05,{ }^{*} \mathrm{p}<0.10$, rob s.e. clustered at school level
Cross sectional weights used
F test of excluded instruments: $F(8,125)=23.25$, Prob $>F=0.000$


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[^1]:    ${ }^{3}$ See Blume, Brock, Durlauf, and Ioannides (2011) for an excellent review of papers on social interactions.

[^2]:    ${ }^{4}$ This research uses data from AddHealth, a program project designed by J. Richard Udry, Peter S. Bearman, and Kathleen Mullan Harris. Special acknowledgment is due to Ronald R. Rindfuss and Barbara Entwisle for assistance in the original design. Persons interested in obtaining data files from AddHealth should contact AddHealth, Carolina Population Center, 123 W. Franklin Street, Chapel Hill, NC 27516-2524, USA (addhealth@unc.edu).

[^3]:    ${ }^{5}$ We do not include the average age of the peers, due to the very high correlation with the age of the respondent (in most cases the respondent and her friends have the same age).

[^4]:    ${ }^{6}$ In the in-school survey adolescents had nominated on average 6 friends. We consider friends that have completed the in-home interview of Wave III in order to have information about their relationship history. Given that only 15,000 out of 90,000 students participated in Wave III the number of peers is reduced substantially.

[^5]:    ${ }^{7}$ Alternatively we tried to use the characteristics of friends of friends who are not directly connected to the respondents but they turned out to be very weak instruments.

[^6]:    ${ }^{8}$ The survey had a separate section about "special" friends and therefore the respondents were not supposed to include them among the nominated friends.

[^7]:    ${ }^{9}$ Another possible mechanism is transmission of information about marriage or cohabitation through peers. However, the absence of spill-over effects from friends of friends (Table 3, column 2) suggests that such a mechanism is unlikely to be at work.
    ${ }^{10}$ These respondents belonged to the youngest cohort of Wave I, this is why their average age and the $\%$ married is lower than those of all the respondents.

[^8]:    ${ }^{11}$ There is a question about the reason why the friendship ended and the most common answer is "it just happened/you drifted apart" among the alternatives: the friend moved away, you moved

[^9]:    ${ }^{1}$ Individuals with non missing relationship history
    ${ }^{2}$ Corrected for survey design

