# Examination Behavior - Gender Differences in Preferences?* 

by

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#### Abstract

A unique examination strategy in first year microeconomics courses is used to test for gender differences in preferences in examination behavior. Students have the possibility of attaining a seminar bonus on the final exam for near-perfect seminar attendance and are given two voluntary initial quizzes during the semester. At the final exam, the scores received on initial quizzes can either be accepted as is, or students can attempt to improve their marks by answering similar quiz questions on the exam. Results suggest that female students are more likely to take initial quizzes and receive a seminar bonus but are less likely to re-take quizquestions on the final exam. These results suggest higher risk aversion among female students relative to male students, a behavioral difference with tangible implications in terms of final grades on the course.


[^0]
## 1. Introduction

At present, there is a large and growing research interest on gender differences in preferences (see Croson and Gneezy (2009) and Bertrand (2010) for excellent overviews of this literature). Women, in this largely experimental literature, have been found to be more riskaverse on average, less willing to compete, less overconfident, more altruistic and more inequality averse than likewise men. There is however a "striking" lack of research in real settings establishing the empirical relevance of these factors for actual outcomes (Bertrand, 2010). We hope to fill this gap in at least one context by studying gender differences in behavior in an academic setting, namely among undergraduate students in a first year economics course at Stockholm University. Students in this course face a number of decisions regarding examination strategy, both prior to and during the actual final examination, with potential repercussions for actual educational outcomes. The aim of our study is to examine if there are gender differences in behavior in examination strategies and to estimate to what degree these systematic behavioral differences are costly in terms of final grades on the exam (and the course).

Students in the first year microeconomic courses at Stockholm University are offered two voluntary quizzes during the semester, the scores of which are credited to the final exam. Quiz takers, however, also have the opportunity of re-taking corresponding quiz questions on the final exam to improve (or lower) their earlier established scores. In addition, students can be awarded a seminar bonus on the final examination for near-perfect seminar attendance during the semester. This set-up allows us to investigate gender differences in examination behavior in three ways. First, taking a quiz is one way to increase the probability of higher exam scores since quiz-takers can redo the quiz on the exam (i.e., have a second chance). Note that quiz-takers also have potentially more time to devote to other exam questions given the fixed time constraint for the final exam. Second, near-perfect seminar attendance yields both an examination bonus and, again, more time on the final exam, both of which should increase the probability of attaining higher final exam scores. Third, quiz-takers have the opportunity of improving their scores by redoing one or both of the quizzes on the exam. The question to answer is whether there are gender differences in examination behavior regarding these three choices and if these behavioral differences are costly in terms of final marks on the exam.

Note that there are a number of potential mechanisms behind observed behavioral
differences in examination strategies having to do with, for example, gender differences in confidence, risk-aversion, competitiveness and/or time usage (such as procrastination) which are difficult to tease out, especially in the type of real world setting examined here. Bertrand (2010) notes, for example, that gender differences in risk attitudes may be a consequence of male overconfidence in their relative ability. Likewise, women may systematically underperform in competitive environments and shy away from competition, both of which may be due to gender differences in confidence or risk attitudes. A thoughtfully carried out experimental study can at times get close to weeding out the influence of a particular type of preference component on outcomes of interest, for example, by changing a controlled environment and introducing competition without a change of risk-levels and/or by controlling for risk via attitudinal survey questions. The question remains, however, if it is possible to distinguish between, for example, risk-aversion and confidence in a real world setting. That said, it is interesting to examine to what degree gender differences in preferences, regardless of its source, matter for real outcomes in real settings.

A related issue concerns other factors that may be considered as part of, or correlated to, individual preferences. If underlying skills (ability, motivation etc.) are correlated with components such as risk aversion and confidence, it is unclear how to interpret gender differences in behavior if such differences are not taken into account. Despite the advantages of experimental studies described above, many do not account for how men and women are selected into any given experiment. This implies that observed gender differences could be due to an underlying difference in skill distribution. Our study analyzes gender differences in preferences as manifested in behavioral choices concerning examination strategies among undergraduate university students. Although we cannot separate between different potential mechanisms behind observed gender differences, we discuss to what degree different mechanisms are consistent with reported results. We do, however, have rich information on students, including grades in the high school math courses that are a pre-requisite for this undergraduate course and can therefore account for any systematic underlying selection in relevant (for the course) skills by gender.

Based on individual information on 2,121 students who completed the first year microeconomics course during the five semesters from 2006 to 2008, female students are found to be more likely to take one or both of the initial quizzes offered during the semester, all else equal, and are also more likely to receive a seminar bonus. During the actual exam,
female quiz-takers are less likely to re-take corresponding quiz questions than male quiztakers. All three strategies (seminar attendance, quiz taking and quiz re-taking) are correlated with higher final exam scores implying that female students win via their higher seminar attendance and higher propensity to take initial quizzes, but lose due to their lower propensity to re-take quiz questions on the final exam.

Results presented in this paper complement the existing literature on gender differences in preferences and may shed some light on how pedagogical methods differentially influence education strategies by gender, with potential repercussions for educational outcomes. Note that the decisions students make concerning exam strategies have private consequences only implying that behavioral differences are not driven by gender differences in care and concern for others.

The rest of the paper is organized as follows. In Section 2 we describe the examination procedure and in Section 3 the data and the empirical set-up. Results are reported in Section 4 followed by a discussion of the results in Section 5. The paper is concluded in Section 6.

## 2. The Exam Procedure

The exam procedure of the first year microeconomics course at Stockholm University provides a great opportunity to study gender differences in examination behavior. All students have the opportunity to take two voluntary quizzes offered during the first and third quarter of the semester. Quizzes aim to test student knowledge and to motivate students to study for exams at an early stage of the course. Both initial quizzes consist of ten multiple-choice questions with a maximum possible score of ten points per quiz.

Students who take one or both quizzes during the semester have the option of skipping corresponding quiz-like questions on the final exam. Original quiz results are then awarded to the corresponding question on the exam. However, quiz-takers also have the option to re-take the corresponding quiz question on the exam in an attempt to improve the original score. If eligible students chose to re-take a quiz question on the exam, the scores received on the exam are final, regardless of whether the student improves his/her score or not. As such, re-taking quizzes on the final exam implies a private risk of lowering quiz scores. Students who did not take the initial quizzes are required to answer the corresponding questions on the final exam or receive zero points for that question. Moreover, students with near-perfect seminar attendance during the course are automatically rewarded ten points on the final exam.

Students who did not attend seminars have the option of answering an additional question on the exam with a maximum possible score of ten points.

This examination procedure provides three ways of testing for gender differences in behavior. The first concerns gender differences in the probability of taking one or both of the voluntary quizzes offered during the course. Taking a voluntary quiz provides an opportunity to improve final exam scores as quiz takers are allowed to redo the corresponding quiz questions on the final exam. The second concerns gender differences in the probability of retaking quiz questions on the final exam among those students that initially took one or both of the quizzes offered during the course. Students who took at least one quiz know their initial scores, can read and evaluate the corresponding questions on the final exam and can thereafter choose whether or not to attempt an improvement by re-doing the quiz questions. The third concerns gender differences in attending seminars as seminar attendance yields a bonus of ten points on the exam.

## 3. Data and Empirical Setup

### 3.1 Data

Data stems from individual information on quizzes and exams taken during first year microeconomics courses at the Department of Economics, Stockholm University, from the fall term 2006 to the spring term 2008. In total, 2,349 students were enrolled in this course during this period and took the final exam. We restrict the analysis to the 2,021 students ( 47 percent of which are female) with background information on the high school math courses that are a pre-requisite for this course. Students at Stockholm University have the possibility of taking the first year microeconomics course via one of nine different academic programs or by registering independently for the course. In our sample, 58 percent of the female students and 56 percent of the male students are enrolled through programs, the remainder independently. ${ }^{1}$

In the empirical analysis two different samples are used. The first sample consists of all students who participated in the course. With this sample we can study gender differences in seminar attendance and the probability of taking one or both of the initial quizzes offered in

[^1]the course. The second sample consists of student-quiz observations based on those students who took at least one quiz. With this sample, we can study gender differences in the probability of re-taking a quiz on the final exam.

Sample means on all students are presented in Table 1, by gender. Male and female students have on average the same results on the final exam, an exam that consists of ten questions with a total potential total score of 100 points. Sample means also indicate that a significantly greater proportion of female students than male students take one or both of the initial quizzes. In addition, a greater proportion of female students than male students attend seminars. Female students are under-represented at the highest high school math level course. Finally, math scores differs somewhat between female and male students.

Table 1. Sample Means Individual observations (standard errors in parentheses)

|  | Female Students | Male Students | Female-Male Difference |
| :--- | :---: | :---: | :---: |
| Birth year | 1983.9 | 1984.1 | -0.23 |
|  | $(0.13)$ | $(0.11)$ | $(0.17)$ |
| Total exam score (0-100) | 62.9 | 62.6 | 0.25 |
|  | $(0.59)$ | $(0.57)$ | $(0.82)$ |
| Wrote at least one quiz | 0.944 | 0.899 | $0.045^{* * *}$ |
|  | $(0.007)$ | $(0.009)$ | $(0.012)$ |
| Wrote two quizzes | 0.835 | 0.787 | $0.048^{* * *}$ |
|  | $(0.012)$ | $(0.012)$ | $(0.017)$ |
| Seminar attendance | 0.876 | 0.812 | $0.063^{* * *}$ |
|  | $(0.010)$ | $(0.012)$ | $(0.016)$ |
| Math level (1-4) |  |  |  |
| Level 1 (lowest) | 0.038 | 0.024 | 0.014 |
|  | $(0.006)$ | $(0.005)$ | $(0.007)$ |
| Level 2 | 0.180 | 0.118 | $0.061^{* * *}$ |
|  | $(0.012)$ | $(0.010)$ | $(0.015)$ |
| Level 3 | 0.526 | 0.428 | $0.098^{* * *}$ |
|  | $(0.016)$ | $(0.015)$ | $0.022)$ |
| Level 4 (highest) | 0.256 | 0.430 | $-0.174^{* * *}$ |
|  | $(0.014)$ | $(0.015)$ | $(0.020)$ |
| Math score (1-4) |  |  |  |
| Fail | 0.084 | 0.110 | $-0.026^{*}$ |
|  | $(0.009)$ | $(0.009)$ | $(0.013)$ |
| Pass (lower level) | 0.467 | 0.535 | $-0.068^{* * *}$ |
|  | $(0.016)$ | $(0.015)$ | $(0.022)$ |
| Pass (higher level) | 0.0334 | 0.244 | $0.091^{* * *}$ |
|  | $(0.015)$ | $(0.013)$ | $(0.020)$ |
| Pass with distinction | 0.114 | 0.111 | 0.003 |
|  | $(0.010)$ | $(0.009)$ | $(0.014)$ |


| No. of individuals | 996 | 1,125 | 2,121 |
| :--- | :--- | :--- | :--- |

Sample means based on student-quiz observations are shown in Table 2. Among quiz takers, approximately 28 percent of both female and male quiz-takers re-take a quiz question on the exam. ${ }^{2}$ It is difficult to interpret this as indication of no behavioral differences between males and females. Rational behavior suggests that the probability of re-taking a quiz should be lower for those with higher initial quiz scores since attempting to improve one's score becomes more difficult with higher initial scores, all else equal. As shown in Table 2, female students have lower average initial quiz scores than male students. The distribution of initial quiz scores by gender is shown in Figure 1. Female students have initial quiz scores of 5 and 6 to a relatively larger extent than male students while male students have initial quiz scores of 9 and 10 to a relatively larger extent than female students. Thus, in the empirical analysis, one should at the very least, compare gender differences in the probability of retaking the quiz conditioning on initial quiz scores.

[^2]Table 2: Sample Means, Quiz Observations (standard errors in parentheses)

|  | Female Students | Male Students | Female-Male <br> Difference |
| :--- | :---: | :---: | :---: |
| Birth Year | 1983.9 | 1984.2 | 0.23 |
|  | $(0.10)$ | $(0.09)$ | $(0.13)$ |
| Total Exam Score (0-100) | 64.5 | 65.1 | 0.61 |
|  | $(0.42)$ | $(0.40)$ | $(0.29)$ |
| Initial Quiz Score | 6.81 | 7.04 | $-0.23 * * *$ |
|  | $(0.046)$ | $(0.046)$ | $(0.066)$ |
| Re-Take Quiz | 0.288 | 0.278 | 0.009 |
|  | $(0.011)$ | $(0.010)$ | $(0.015)$ |
| Math level (1-4) |  |  |  |
| Level 1 (lowest) | 0.035 | 0.025 | 0.010 |
|  | $(0.004)$ | $(0.004)$ | $(0.006)$ |
| Level 2 | 0.179 | 0.123 | $0.056^{* * *}$ |
|  | $(0.009)$ | $(0.008)$ | $(0.012)$ |
| Level 3 | 0.527 | 0.431 | $0.096^{* * *}$ |
|  | $(0.012)$ | $(0.011)$ | $(0.016)$ |
| Level 4 (highest) | 0.259 | 0.421 | $-0.162^{* * *}$ |
|  | $(0.010)$ | $(0.011)$ | $(0.015)$ |
| Math score (1-4) |  |  |  |
| Fail | 0.082 | 0.104 | $-0.022^{* *}$ |
|  | $(0.007)$ | $(0.007)$ | $(0.010)$ |
| Pass (lower level) | 0.464 | 0.541 | $-0.077^{* * *}$ |
|  | $(0.012)$ | $(0.011)$ | $(0.016)$ |
| Pass (higher level) | 0.334 | 0.245 | $0.089^{* * *}$ |
|  | $(0.011)$ | $(0.010)$ | $(0.015)$ |
| Pass with distinction | 0.120 | 0.110 | 0.009 |
|  | $(0.008)$ | $(0.007)$ | $(0.011)$ |
| No. of Quiz Observations | 1772 | 1896 | 3668 |
| No. of Individuals | 940 | 1,011 | 1,951 |

Figure 1. Distribution of initial quiz scores, by gender.


### 3.2 Empirical Setup

To test for systematic gender differences in the propensity to take one or both of the initial quizzes offered during the course, we estimate the following linear probability models:

$$
\begin{align*}
& \text { At least one quiz } z_{i}=\text { female }_{i}+\mu_{s}+\gamma_{c}+\lambda_{b}+\kappa_{g}+\kappa_{l}+\delta_{q}+\varepsilon_{i}  \tag{1}\\
& \text { Two quizzes }_{i}=\text { ffemale }_{i}+\mu_{s}+\gamma_{c}+\lambda_{b}+\kappa_{g}+\kappa_{l}+\delta_{q}+\varepsilon_{i} \tag{2}
\end{align*}
$$

where At least one quiz is a dummy variable equal to one if student $i$ took at least one of the initial voluntary quizzes offered during the course and zero otherwise. Correspondingly, two quizzes is a dummy variable equal to one if student $i$ took both of the initial quizzes and zero otherwise. The variable of interest is the female dummy variable which takes the value one if student $i$ is a female and zero otherwise and $\beta$ measures the average differences in behavior between females and males. The remaining parameters represent a full set of controls for semester effects ( $\mu_{s}, s=1, \ldots, 5$ ), course code effects ( $\gamma_{c}, c=1, \ldots, 10$ ), birth year effects ( $\lambda_{b}$, $b=1952, \ldots, 1990$ ), quiz question effects ( $\delta_{q}, q=1,2$ ), as well as high school math grades ( $\kappa_{\phi}$, $g=1, ., ., 4)$ and math level ( $\kappa_{l}, l=1, \ldots, 5$ ) effects.

An analogous regression for the probability of getting a seminar attendance bonus (10
points) on the exam due to near-perfect seminar attendance during the semester of study is also estimated. The equation is specified in the same way as equations (1) and (2) above except that no control for quiz question is included.

To test for gender differences in the propensity to re-take quiz questions on the exam, the following linear probability model, based on student-quiz level data, is estimated:

$$
\begin{equation*}
\text { Retake quiz }_{\text {iq }}=\text { Bfemale }_{i}+\theta_{\text {quiz_score }+\mu_{s}+\gamma_{c}+\lambda_{b}+\kappa_{g}+\kappa_{l}+\delta_{q}+\alpha_{a}+\varepsilon_{i q},} \tag{3}
\end{equation*}
$$

where Retake quiz is a dummy variable equal to one if student $i$ re-takes the corresponding quiz question, $q$, on the exam and zero otherwise. Over and beyond the controls described above, estimation of re-take propensities includes a full set of controls for initial quiz scores $\theta_{\text {quiz_score }}$ (quiz_score $=0, \ldots, 10$ ) and seminar attendance, $\alpha_{a}$. The latter is important as those without the seminar bonus must answer an additional 10 point question on the exam implying differences in time constraints between these students and those with the bonus which may influence re-take propensities.

The ultimate goal with all models described above is to capture gender differences in behavior that do not stem from any other non-observable gender differences, such as underlying ability. The strategy used is inevitably based on selection on observables. We argue, however, that we control for the most crucial factors possible given the setting. By using information on math grades in high school we can be reasonably certain that our estimated gender differences in behavior do not reflect gender differences in the underlying skills relevant for the course in question. Moreover, entrance into a given program within this economics course is competitive and course code effects should capture any systematic differences in entrance requirements.

## 4. Results

### 4.1. Gender differences in the probability of taking initial quizzes

Results based on estimations of equations (1) and (2) are reported in Table 3. In Column 1 the dependent variable is whether the student took at least one quiz and in Column 2 the dependent variable is whether the student took two quizzes. Coefficient estimates show that female students are associated with a 3.8 percentage point higher probability of taking quizzes than male students. These effects correspond to a 4.1 (4.7) percent higher probability for
female students to take at least one quiz (both quizzes). Gender differences in the propensity to attend seminars are also estimated and results reported in in Column 3 (Table 3). Results indicate that female students are more likely to attend seminars and therefore receive the seminar bonus on the exam than likewise male students.

Table 3. Gender Differences in the probability of taking a quiz and of attending seminars

|  | The Probability of <br> Taking At Least One <br> Quiz: | The Probability of <br> Taking Two <br> Quizres: | The Probability of <br> Getting a Seminar <br> Attendance Bonus: |
| :--- | :---: | :---: | :---: |
| Female | $0.038^{* * *}$ | $0.038^{* *}$ | $0.055^{* * *}$ |
| Take quiz(zes) <br> baseline |  |  |  |
| Percent effect | 0.920 | $0.017)$ | 0.810 |
| No. of <br> Observations | 0.041 | 0.047 | 0.0642 |

Note: Linear probability models on dummy variables indicating whether or not individuals took at least one quiz or both quizzes and if they attended seminars and received seminar bonus. Control for semester (1-5), course code (1-10), birth year (1952-1990), math level (1-5) and math grade (1-4) including full interaction between math level and math grade.

### 4.2 Gender differences in the probability of re-taking quizzes on the exam

Results from estimations on the probability to re-take quizzes on the exam, i.e., Equation (3), are reported in Table 4. Results show that female students are significantly less likely than male students to re-take quiz questions on the exam. In percentage terms, female students are almost 10 percent less likely to re-take a quiz on the final exam than male students, all else equal.

Table 4. Gender differences in the probability of re-taking quizzes on the final exam (quiz observations)

| Female | Re-take Propensity |
| :--- | :---: |
|  | $-0.028^{* *}$ |
| Re-take (quiz) baseline | $(0.013)$ |
| Percent effect | 0.283 |
| No. of Observations | -0.097 |
| N | 3,668 |

Note: Linear probability models on dummy variables indicating whether or not individuals re-took at least one quiz. Control for quiz question (1-2), seminar attendance, semester (1-5), course code (1-10), birth year (19521990), math level (1-5) and math grade (1-4) including full interaction between math level and math grade and quiz question and initial quiz score including full interaction between quiz question and initial quiz score. Robust standard errors in parentheses.

Average difference can mask differences across the skills distribution. There are a number of studies that suggest that women who choose to compete, especially in male dominated environments, are as competitive as the men in these environments. ${ }^{3}$ Estimation of gender differences in re-take propensities across the distribution of initial quiz scores (not shown) shows that results are driven by those with initial quiz scores of four and five. Re-take propensities are similar for female and male students with higher scores confirming previous results that indicate smaller gender differences in behavior at the high end of the skills distribution.

### 4.3 Potential costs of gender differences in exam behavior

Before discussing potential explanations behind observed gender differences in examination behavior, let us first look at how these behavioral differences (quiz-taking, quiz re-taking and seminar attendance) correlate with other examination outcomes. The difference between quiz-takers and non-quiz-takers in final exam scores is shown in Table 5. Quiz taking is associated with higher final exam scores. Estimation controlling for all relevant covariates yields results showing that taking at least one quiz (two quizzes) is associated with approximately 12 higher points on the final exam in comparison to not taking any quizzes (which correspond to 17-20 percent higher final exam scores).

[^3]Table 5. The Correlation between Quiz-Taking and Final Exam Scores

|  | (1) | (2) | (3) | (4) |
| :---: | :---: | :---: | :---: | :---: |
| At Least One Quiz | $\begin{gathered} 12.1^{* * *} \\ (1.61) \end{gathered}$ | $\begin{gathered} 11.8^{* * *} \\ (1.87) \end{gathered}$ | -- | -- |
| Two Quizzes | -- | -- | $\begin{gathered} 10.9^{* * *} \\ (1.00) \\ \hline \end{gathered}$ | $\begin{gathered} 12.0^{* * *} \\ (1.28) \\ \hline \end{gathered}$ |
| Female | $\begin{gathered} -1.61 \\ (0.74) \\ \hline \end{gathered}$ | $\begin{aligned} & -2.40 \\ & (3.26) \end{aligned}$ | $\begin{gathered} -1.60^{* *} \\ (0.73) \\ \hline \end{gathered}$ | $\begin{aligned} & \hline 0.48 \\ & (1.83 \end{aligned}$ |
| Female* At Least One Quiz | -- | $\begin{gathered} \hline 0.85 \\ (3.26) \\ \hline \end{gathered}$ | -- | -- |
| Female* Two Quizzes | -- | -- | -- | $\begin{gathered} -2.54 \\ (1.96) \\ \hline \end{gathered}$ |
| Final exam baseline | 62.7 |  |  |  |
| Percent effect | 19.5 | 19.0 | 17.4 | 19.1 |
| No. of Observations | 2,121 |  |  |  |

Note: OLS estimation of exam scores $(0-100)$ on quiz taking, defined as dummy variables equal to one if an individual has taken at least one quiz, alternatively both quizzes, and zero otherwise. Control for quiz question (1-2), seminar attendance, semester (1-5), course code (1-10), birth year (1952-1990), math level (1-5) and math grade (1-4) including full interaction between math level and math grade. Robust standard errors in parentheses.

Thus, quiz takers do better on final exams, either due to a systematically different type of study behavior which voluntary quizzes give rise to (including the possibility of re-taking quiz questions on the final exam) or because quiz-takers are selected among the more ambitious and scholarly students enrolled in the course. As we control for differences in relevant math skills, more weight should perhaps be given to the first explanation. Female quiz takers who take at least one initial quiz do not appear to differ from likewise male quiz takers in terms of the quiz premium on final exams as indicated by the insignificant coefficient on the genderquiz interaction terms (Columns 2). The quiz premium for those who take two initial quizzes appears to differ somewhat between female and male students (Columns 4). The difference is not however significant.

Turning instead to an examination of the correlation between re-taking quizzes on the exam and final exam scores (based on quiz observations), we see from results reported in Table 6 that quiz re-takers are associated with a significantly higher final quiz scores in comparison with those that abstain from re-taking the quiz. Quiz re-takers are associated with, on average, 1.3 higher points than those that do not re-take the quiz on the final exam. No gender differences in quiz improvement are noted as shown by the insignificant estimated coefficient on the interaction between gender and a re-take dummy variable.

Table 6. The correlation between quiz re-taking and final exam scores (quiz observations)

|  | Quiz Score: |  |
| :---: | :---: | :---: |
| Quiz Re-Take | $\begin{gathered} 1.266 * * * \\ (0.063) \\ \hline \end{gathered}$ | $\begin{gathered} 1.242 * * * \\ (0.083) \\ \hline \end{gathered}$ |
| Female | $\begin{aligned} & \hline-0.015 \\ & (0.033) \\ & \hline \end{aligned}$ | $\begin{gathered} -0.028 * * \\ (0.013) \\ \hline \end{gathered}$ |
| Quiz Re-Take* Female | -- | $\begin{gathered} \hline 0.046 \\ (0.112) \\ \hline \end{gathered}$ |
|  | Other Exam Question Scores: |  |
| Quiz Re-Take | $\begin{gathered} \hline 6.18^{* * *} \\ (0.659) \end{gathered}$ | $\begin{gathered} 7.00^{* * *} \\ (0.81) \end{gathered}$ |
| Female | $\begin{gathered} 0.44 \\ (0.536) \\ \hline \end{gathered}$ | $\begin{gathered} 0.91 \\ (0.62) \\ \hline \end{gathered}$ |
| Quiz Re-Take* Female | -- | $\begin{gathered} -1.63 \\ (1.01) \\ \hline \end{gathered}$ |
| No. of Observations | 3,668 | 3,668 |

Note: OLS estimation of quiz scores ( $0-10$ ) and exam scores ( $0-100$ ) on quiz re-taking, defined as dummy variable equal to one if an individual that has taken at least one of the initial voluntary quizzes and re-taken it on the final exam. All estimation controls for quiz question (1-2), seminar attendance, semester of study (1-5), course code (1-10), birth year (1952-1990), math level (1-5) and math grade (1-4) as well as a full interaction between math level and math grade. Robust standard errors reported in parenthesis.

Re-taking quizzes on the exam implies a time investment. Students must weigh in to what degree they will benefit from re-taking quiz questions given that less time can be invested in the remaining mandatory questions on the exam. Students are allotted a maximum five hours to complete final exams. Abstaining from re-taking quiz questions on the exam may therefore be a rational strategy for students who feel that their time is better invested in answering other final exam questions. If such is the case, we would expect that non re-takers have higher scores on the other exam questions than re-takers, especially for female students. Results in the second panel of Table 6, however, show that re-takers are associated with higher points (a 6-7 point advantage) on the remaining exam questions and that there is no gender difference in this association. These correlations therefore suggest that it is not time constraints driving female quiz takers to abstain from re-taking quizzes on the exam, rather there appears to be spill-over effects on other exam questions implying that the selection of students who prepare for the possibility of re-taking quizzes on the exam improve their scores over and beyond the improvement on the quiz questions alone. Again no gender differences in this correlation are found.

Taken together, results show that quiz re-takers are associated with higher quiz scores
on the exam as well as higher marks on other exam questions. As such, quiz re-taking is also associated with an approximate eight point advantage in terms of higher total final exam scores.

One way of measuring the implications of observed gender differences in examination behavior is to estimate how final exam scores would be affected if female students behaved like male students. In order to do so, we estimate the following equation on final exam scores based on the full sample of students: ${ }^{4}$

Final Exam Score ${ }_{i}=\beta_{l}$ Seminar $+\beta_{2}$ Quiz $+\beta_{3}$ Re-take $+\theta_{\text {quiz_score }}+\mu_{s}+\gamma_{c}+\lambda_{b}+\kappa_{g}+\kappa_{l}+\delta_{q}+\varepsilon_{i}$

Table 7. Predicted Final Exam Scores if Female Students Behave like Male Students

|  | Coefificient (Robust St. <br> Error) | Predicted Final Exam Score if Female <br> Students behaved like Male Students <br> (deviation from baseline) |
| :--- | :---: | :---: |
| Seminar | $12.67^{* * *}(0.98)$ | $61.89(-0.84)$ |
| Quiz | $4.81^{* * *}(2.12)$ | $62.46(-0.26)$ |
| Re-take | $8.14^{* * *}(0.91)$ | $63.59(+0.86)$ |
| Baseline prediction for females |  | 62.73 |

Thereafter, we predict female exam scores based on male examination behavior. In other words, given that the coefficient estimate indicates that seminar attendance is associated with 12.7 higher points on the final exam, if female students behaved as male students in terms of seminar attendance, they would on the margin, be penalized by a lower final exam score of 0.84 points. Likewise, if female students behaved like male students in terms of taking initial quizzes, they would again be penalized by lower final exam scores ( 0.26 points lower). On the other hand, if female students re-took quizzes on the exam to the same extent as male students, they would gain 0.86 points on the final exam (on the margin). ${ }^{5}$ Female students therefore win through their higher propensities to attend seminars and take quizzes but lose through their lower propensity to re-take quiz questions on the exam.

[^4]
## 5. Potential Explanations for Gender Differences in Examination Behavior

There are a number of potential mechanisms behind our observed behavioral differences in examination strategies having to do with, for example, gender differences in confidence, riskaversion, competitiveness and/or time usage (such as procrastination), which are difficult to tease out in the type of real world setting examined here. Below we discuss the credibility of some of these mechanisms in explaining observed results.

Seminar attendance and quiz taking are two ways of gaining more time for other questions during the final exam. The seminar bonus automatically yielded the highest possible scores on one of the final exam questions and students that were satisfied with initial quiz scores could refrain from re-taking corresponding quiz questions on the final exam and devote this time to other questions. Taking one or both of the initial quizzes also provided students with a second chance to improve their scores on the final exam. As such, seminar attendance and quiz taking must be considered risk averse strategies in examination behavior. Our results showing higher propensities among female students to take initial quizzes and attain a seminar bonus are therefore consistent with women being, on average, more risk avert than men.

Re-taking a quiz on the final exam is associated with a private risk of lowering initial quiz scores. Although observed gender differences in re-taking propensities are driven by students with relatively low initial quiz scores, such as 5 out of 10 , there is still a risk involved in attempting to improve initial scores. Our result that female students are less inclined to retake the quizzes, all else equal, is therefore also consistent with an interpretation that female students are more risk averse than male students. Note that when analysing gender differences in re-take propensities, we use a selection of students who took the initial quizzes. Since female students took quizzes to a larger extent, and if taking a quiz is associated with a higher degree of risk aversion, it must be the case that we have stronger selection of risk averse males students compared to female students among initial quiz-takers. As such, it is likely that we underestimate gender differences in risk aversion as measured by re-take propensities.

An alternative interpretation of our results is that male students are more confident than female students. Overconfidence is a relative concept. Men could be confident (and correct) in their interpretation of their abilities whereas women are under-confident, or men could be overconfident (i.e., have an incorrect estimation of their ability) whereas women have a correct estimation of their ability. Irrespective of whether male students are overconfident or female under-confident, given that male students have the same ambition as female students
to maximize final exam scores, our results may reflect male perceptions of a higher capability to receive high exam scores without the bonus, in terms of time or scores, of taking an initial quiz and/or attending seminars. Higher confidence, on average, can also explain why male students re-take quizzes to a large extent, all else equal.

The exam procedure under study here did not involve any direct rewards that depend on students' relative performance implying that gender differences in competitive behavior is unlikely to be a relevant explanation for presented results. There is one other potential explanation which concerns gender differences in time-use. Male students may be more prone to procrastinate implying that they study less at the beginning of the course and more intensely close to the date of the final exam. Gender differences in procrastination may then explain why male students take initial quizzes to a smaller extent and have lower seminar attendance. According to course evaluations of students who attended the microeconomics course in 2011, a larger fraction of female than male students spend more than 30 hours a week studying for this course. This does not imply that male students procrastinate to a larger extent than female students, but may be consistent with a story where male students spend less time on their studies and, perhaps, distribute this time differently than female students. Higher male procrastination is also consistent with higher male confidence about doing well on final exams. However, procrastination cannot explain why male students, given than they took an initial quiz, re-take these on the final exam to a larger extent than female students.

Taken together, the observed gender differences in examination behavior found in this study seem consistent with conclusions from the experimental literature that women are more risk averse and/or less (over)confident, on average, than men. Studying gender differences in a natural setting allows us to provide tentative conclusions on how gender differences in preferences affect real outcomes. Attending seminars, taking initial quizzes and re-taking quizzes on the final exam are all correlated with higher final exam scores. As such, the risk averse behavior of female students in terms of higher seminar attendance and taking initial questions, seems to be a better strategy than that followed by male students. On the other hand, female students could improve their results on the final exam by behaving more like men when it comes to re-taking quizzes on the final exam.

## 6. Conclusion

This study has used information on examination behavior among first year microeconomics students at Stockholm University to analyze gender differences in preferences in terms of examination behavior. Students in this course have the possibility of attaining a seminar bonus on the final exam for near-perfect seminar attendance and are given two voluntary initial quizzes during the semester. At the final exam, the scores received on initial quizzes can either be accepted as is, or students can attempt to improve their marks by answering similar quiz questions on the exam. Results suggest that female students are more likely to take initial quizzes and receive a seminar bonus but are less likely to re-take quiz-questions on the final exam. Attending seminars, taking initial quizzes and re-taking quizzes on the final exam are all correlated with higher final exam scores. As such, the risk averse behavior of female students in terms of higher seminar attendance and taking initial questions, seems to be a better strategy than that followed by male students. On the other hand, female students could improve their results on the final exam by behaving more like men when it comes to re-taking quizzes on the final exam.

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[^1]:    ${ }^{1}$ The academic programs are business and administration, accounting, retailing, mathematics and economics, economics and statistics, social planning, economics and political science and social science educators. The largest proportion of students enrolled via an academic program is in a business school program.

[^2]:    ${ }^{2}$ Among students who took both initial quizzes, 25 percent of both male and female students retake one quiz question and 14 percent chose to retake both quiz questions.

[^3]:    ${ }^{3}$ See for example: Master and Meier (1988), Birley (1989), Johnson and Powell (1994), Dwyer et al. (2002), Atkinson et al. (2003), Datta Gupta et al. (2005), Nekby et al. (2008) and Garratt et al. (2011).

[^4]:    ${ }^{4}$ The dummy for re-take is re-coded to be equal to zero for all students that (1) did not re-take a quiz question on the exam (among quiz-takers) and (2) for those students that did not take either of the original quizzes.
    ${ }^{5}$ Estimation of mean differences in re-take behavior by gender is evaluated at an initial quiz score equal to five.

